

A Unified Confidence Sequence for Generalized Linear Models, with Applications to Bandits

Junghyun Lee¹, Se-Young Yun¹, and Kwang-Sung Jun²

¹ Kim Jaechul Graduate School of AI, KAIST, ² Department of Computer Science, University of Arizona
 {jh_lee00, yunseyoung}@kaist.ac.kr, kjun@cs.arizona.edu



Contributions

- A *unified*, state-of-the-art construction of likelihood ratio-based confidence sequence (CS) for any convex generalized linear models (GLMs), with explicit constants!
- A new CS-based algorithm (OFUGLB) that achieves the state-of-the-art regret for self-concordant GLBs.
- Numerical verifications in logistic bandits show the tightness of our new CS and that OFUGLB achieves the best numerical regret by a large margin.

Problem Settings

Generalized Linear Models (GLMs)

For a covariate $\mathbf{x} \in X$ and an *unknown* parameter $\boldsymbol{\theta}_* \in \Theta$, the reward r follows the **GLM** if

$$dp(r|\mathbf{x}; \boldsymbol{\theta}_*) = \exp\left(\frac{r\langle \mathbf{x}, \boldsymbol{\theta}_* \rangle - m(\langle \mathbf{x}, \boldsymbol{\theta}_* \rangle)}{g(\tau)} + h(r, \tau)\right) d\nu, \quad (1)$$

where τ is some known scaling (temperature) parameter, and ν is some known base measure (e.g., Lebesgue, counting).

Assumptions:

- **Assumption 1.** $X \subseteq \mathcal{B}^d(1)$.
- **Assumption 2.** $\boldsymbol{\theta}_* \in \Theta \subseteq \mathcal{B}^d(S) := \{\boldsymbol{\theta} \in \mathbb{R}^d : \|\boldsymbol{\theta}\|_2 \leq S\}$ for some known $S > 0$. Also, Θ is nonempty, compact, and convex with intrinsic dimension d .
- **Assumption 3.** m is three times differentiable and convex, i.e., m''' exists and $\dot{m} := m'' \geq 0$.

Well-known properties:

- **Property 1.** $\mathbb{E}[r|\mathbf{x}, \boldsymbol{\theta}_*] = m'(\langle \mathbf{x}, \boldsymbol{\theta}_* \rangle) \triangleq \mu(\langle \mathbf{x}, \boldsymbol{\theta}_* \rangle)$
 - **Property 2.** $\text{Var}[r|\mathbf{x}, \boldsymbol{\theta}_*] = g(\tau)\dot{m}(\langle \mathbf{x}, \boldsymbol{\theta}_* \rangle)$.
- μ is the **inverse link (mean) function**.

Question #1

Given a (possibly adaptively-collected) sequential data $\{(x_t, r_t)\}_{t \geq 1}$ sampled from any GLM, output the tightest **confidence sequence (CS)** for $\boldsymbol{\theta}_*$, i.e., for any $\delta \in (0, 1)$, $\{\mathcal{C}_t(\delta)\}_{t \geq 1}$ such that $\mathbb{P}[\exists t \geq 1 : \boldsymbol{\theta}_* \notin \mathcal{C}_t(\delta)] \leq \delta$.

Generalized Linear Bandits (GLBs)

First proposed in Filippi et al. [2010] as a nonlinear generalization of linear bandits.

For $t \in [T]$:

- 1 The learner observes a potentially infinite (contextual) arm-set $\mathcal{X}_t \subset X$
- 2 The learner chooses $\mathbf{x}_t \in \mathcal{X}_t$ according to some policy
- 3 Receive a reward $r_t|\mathbf{x}_t \sim p(\cdot|\mathbf{x}_t, \boldsymbol{\theta}_*)$ (Eqn. (1))

Goal. Minimize:

$$\text{Reg}^B(T) := \sum_{t=1}^T \{\mu(\langle \mathbf{x}_{t,*}, \boldsymbol{\theta}_* \rangle) - \mu(\langle \mathbf{x}_t, \boldsymbol{\theta}_* \rangle)\},$$

where $\mathbf{x}_{t,*} := \arg \max_{\mathbf{x} \in \mathcal{X}_t} \mu(\langle \mathbf{x}, \boldsymbol{\theta}_* \rangle)$.

Applications. News recommendations (Bernoulli), social network influence maximization (Poisson), etc [Filippi et al., 2010].

We define the following problem-dependent quantities:

$$\kappa_*(T) := \left(\frac{1}{T} \sum_{t=1}^T \dot{\mu}(\mathbf{x}_{t,*}^\top \boldsymbol{\theta}_*)\right)^{-1}, \quad \kappa_{\mathcal{X}}(T) := \max_{t \in [T]} \max_{\mathbf{x} \in \mathcal{X}_t} \frac{1}{\dot{\mu}(\mathbf{x}^\top \boldsymbol{\theta}_*)},$$

and $\kappa(T) := \max_{t \in [T]} \max_{\mathbf{x} \in \mathcal{X}_t} \max_{\boldsymbol{\theta} \in \Theta} \frac{1}{\dot{\mu}(\mathbf{x}^\top \boldsymbol{\theta})}$.

These can scale *exponentially* in S (e.g., Bernoulli)!

$d\sqrt{T/\kappa_*(T)}$ -type regret has been obtained for bounded GLBs in a concurrent work of Sawarni et al. [2024], but they make use of explicit warmup and consider limited adaptivity setting.

Question #2

Using our tight CS, how do we obtain tight regret bounds for a wide range of GLBs via a *purely optimistic approach*?

A Unified CS for GLMs

We consider log-likelihood-based confidence set “centered” at the *norm-constrained*, batch maximum likelihood estimator (MLE):

$$\mathcal{C}_t(\delta) := \left\{ \boldsymbol{\theta} \in \Theta : \mathcal{L}_t(\boldsymbol{\theta}) - \mathcal{L}_t(\hat{\boldsymbol{\theta}}_t) \leq \beta_t(\delta)^2 \right\}, \quad (2)$$

where $\beta_t(\delta)^2$ is the “radius” of the CS that we will define later, and $\mathcal{L}_t(\boldsymbol{\theta})$ is the negative log-likelihood of $\boldsymbol{\theta}$ w.r.t. data collected up to $t-1$, and

$$\mathcal{L}_t(\boldsymbol{\theta}) := \sum_{s=1}^{t-1} \left\{ \ell_s(\boldsymbol{\theta}) \triangleq \frac{-r_s \langle \mathbf{x}_s, \boldsymbol{\theta} \rangle + m(\langle \mathbf{x}_s, \boldsymbol{\theta} \rangle)}{g(\tau)} \right\}, \quad (3)$$

$$\hat{\boldsymbol{\theta}}_t := \arg \min_{\boldsymbol{\theta} \in \Theta} \mathcal{L}_t(\boldsymbol{\theta}). \quad (4)$$

Theorem 3.1. Let $L_t := \max_{\boldsymbol{\theta} \in \Theta} \|\nabla \mathcal{L}_t(\boldsymbol{\theta})\|_2$ be the Lipschitz constant of $\mathcal{L}_t(\cdot)$ that may depend on $\{(\mathbf{x}_s, r_s)\}_{s=1}^{t-1}$. Then, we have $\mathbb{P}[\exists t \geq 1 : \boldsymbol{\theta}_* \notin \mathcal{C}_t(\delta)] \leq \delta$, where

$$\mathcal{C}_t(\delta) = \left\{ \boldsymbol{\theta} \in \Theta : \mathcal{L}_t(\boldsymbol{\theta}) - \mathcal{L}_t(\hat{\boldsymbol{\theta}}_t) \leq \beta_t(\delta)^2 \right\}, \quad (5)$$

where $\beta_t(\delta)^2 \leq \log \frac{1}{\delta} + d \log \left(e \vee \frac{2eSL_t}{d} \right)$.

For **Bernoulli**, our radius of $\mathcal{O}_\delta(d \log(St/d))$ this is a strict improvement over prior $\mathcal{O}_\delta(d \log(St/d) + S)$ of Lee et al. [2024].

→ **Remark.** This resolves an open problem posed by Lee et al. [2024] on poly(S)-free CS for **Bernoulli**.

We consider the following additional assumption on the GLM:

→ **Assumption 4. (self-concordance)** For some $R_s \in (0, \infty)$, $|\dot{\mu}(\langle \mathbf{x}, \boldsymbol{\theta} \rangle)| \leq R_s \dot{\mu}(\langle \mathbf{x}, \boldsymbol{\theta} \rangle)$ for all $\mathbf{x} \in X, \boldsymbol{\theta} \in \Theta$.

For this class of GLMs, we have a slightly relaxed *ellipsoidal* CS:

Theorem 3.2. With the same notations as Theorem 3.1, we have $\mathbb{P}[\exists t \geq 1 : \boldsymbol{\theta}_* \notin \mathcal{E}_t(\delta)] \leq \delta$, where

$$\mathcal{E}_t(\delta) := \left\{ \boldsymbol{\theta} \in \Theta : \left\| \boldsymbol{\theta} - \hat{\boldsymbol{\theta}}_t \right\|_{\nabla^2 \mathcal{L}_t(\hat{\boldsymbol{\theta}}_t) + \frac{1+SR_s}{2S^2} \mathbf{I}_d} \leq \gamma_t(\delta)^2 \right\}, \quad (6)$$

where $\gamma_t(\delta)^2 := 2(1 + SR_s)(1 + \beta_t(\delta)^2)$.

→ **Remark.** This is easier to implement in practice, and for bandits, this amounts to a closed-form bonus in UCB.

Proof via PAC-Bayes with Uniform Prior/Posterior

1. PAC-Bayesian Time-Uniform Bound.

Lemma 3.3. For any data-independent prior \mathbb{Q} and any sequence of adapted posterior distributions $\{\mathbb{P}_t\}$, the following holds: for any $\delta \in (0, 1)$,

$$\mathbb{P} \left(\exists t \geq 1 : \mathcal{L}_t(\boldsymbol{\theta}_*) - \mathbb{E}_{\boldsymbol{\theta} \sim \mathbb{P}_t} [\mathcal{L}_t(\boldsymbol{\theta})] \geq \log \frac{1}{\delta} + \text{KL}(\mathbb{P}_t || \mathbb{Q}) \right) \leq \delta. \quad (7)$$

Proof sketch. This is a standard recipe using Ville’s inequality and Donsker-Varadhan variational representation of KL; see Chugg et al. [2023] for relevant references.

2. Novel choice of \mathbb{Q} and \mathbb{P}_t .

For $c \in (0, 1]$ to be determined later, we set

$$\mathbb{Q} = \text{Unif}(\Theta), \quad \mathbb{P}_t = \text{Unif}(\tilde{\Theta}_t \triangleq (1-c)\hat{\boldsymbol{\theta}}_t + c\Theta), \quad (8)$$

where $\mathbf{a} + \Theta = \{\mathbf{a} + \boldsymbol{\theta} : \boldsymbol{\theta} \in \Theta\}$ for a vector $\mathbf{a} \in \mathbb{R}^d$.

Then, we have

$$\text{KL}(\mathbb{P}_t || \mathbb{Q}) = \log \frac{\text{vol}(\Theta)}{\text{vol}(\tilde{\Theta}_t)} = d \log \frac{1}{c}.$$

3. Lipschitzness of $\mathcal{L}_t(\cdot)$.

We also have that

$$\mathbb{E}_{\boldsymbol{\theta} \sim \mathbb{P}_t} [\mathcal{L}_t(\boldsymbol{\theta})] = \mathcal{L}_t(\hat{\boldsymbol{\theta}}_t) + \mathbb{E}_{\boldsymbol{\theta} \sim \mathbb{P}_t} [\mathcal{L}_t(\boldsymbol{\theta}) - \mathcal{L}_t(\hat{\boldsymbol{\theta}}_t)] \leq \mathcal{L}_t(\hat{\boldsymbol{\theta}}_t) + 2SL_t c,$$

where the last inequality follows from the Lipschitzness of $\mathcal{L}_t(\cdot)$ and the observation that for $\boldsymbol{\theta} = (1-c)\hat{\boldsymbol{\theta}}_t + c\tilde{\boldsymbol{\theta}} \in \tilde{\Theta}_t$ and $\|\boldsymbol{\theta} - \hat{\boldsymbol{\theta}}_t\|_2 = c\|\tilde{\boldsymbol{\theta}} - \hat{\boldsymbol{\theta}}_t\|_2 \leq 2Sc$. We conclude by choosing minimizing over $c \in (0, 1]$. The expression in Theorem 3.1 follows from $c = 1 \wedge \frac{d}{2SL_t}$. □

→ **Remark.** Such choices of \mathbb{Q} and \mathbb{P}_t have been considered previously in universal portfolios [Blum and Kalai, 1999] and fast rates in online learning [Foster et al., 2018]. This is the first time such a translated/shrunk posterior has been used in the PAC-Bayes context.

OFUGLB

OFUGLB is of the following form:

- 1 Obtain $\hat{\boldsymbol{\theta}}_t$ (Eqn. (4)) and $\mathcal{C}_t(\delta)$ (Theorem 3.1)
- 2 Solve $(\mathbf{x}_t, \boldsymbol{\theta}_t) = \arg \max_{\mathbf{x} \in \mathcal{X}_t, \boldsymbol{\theta} \in \mathcal{C}_t(\delta)} \mu(\langle \mathbf{x}, \boldsymbol{\theta} \rangle)$
- 3 Play \mathbf{x}_t , then observe/receive a reward $r_t \in \{0, 1\}$.

We then have the following *state-of-the-art* regret bound:

Theorem 4.1. OFUGLB attains the following regret bound with probability at least $1 - \delta$:

$$\text{Reg}^B(T) \lesssim_\delta d \sqrt{\frac{g(\tau)T}{\kappa_*(T)}} + d^2 R_S R_\mu \sqrt{g(\tau) \kappa(T)},$$

where $R_\mu := \max_{\mathbf{x} \in \mathcal{X}_{[T]}, \boldsymbol{\theta} \in \Theta} \dot{\mu}(\langle \mathbf{x}, \boldsymbol{\theta} \rangle)$.

→ **Remark.** Nontrivial technical contributions, including a new optimistic upper bound of regret, self-concordant control, etc.

Linear bandits. $\tilde{\mathcal{O}}(\sigma d \sqrt{T})$

→ matches prior state-of-the-art [Flynn et al., 2023]

Logistic bandits. $\tilde{\mathcal{O}}(d \sqrt{T/\kappa_*(T)} + d^2 \kappa(T))$

→ first poly(S)-free regret with *purely optimistic approach*, improves upon OFULog+ of Lee et al. [2024]!

Poisson bandits. $\tilde{\mathcal{O}}(dS \sqrt{T/\kappa_*(T)} + d^2 e^{2S} \kappa(T))$

→ first regret guarantee!

Experiments for Logistic Bandits

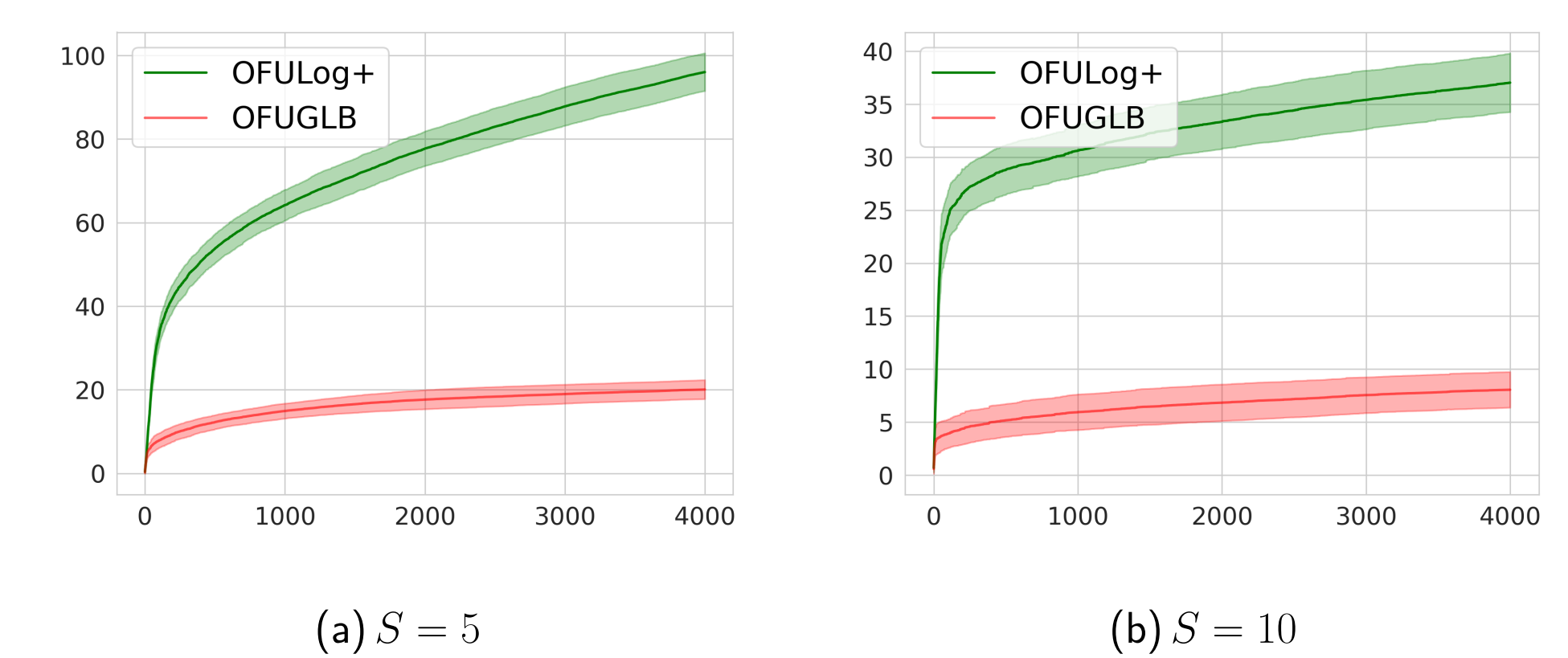


Figure 1: Numerical regrets.

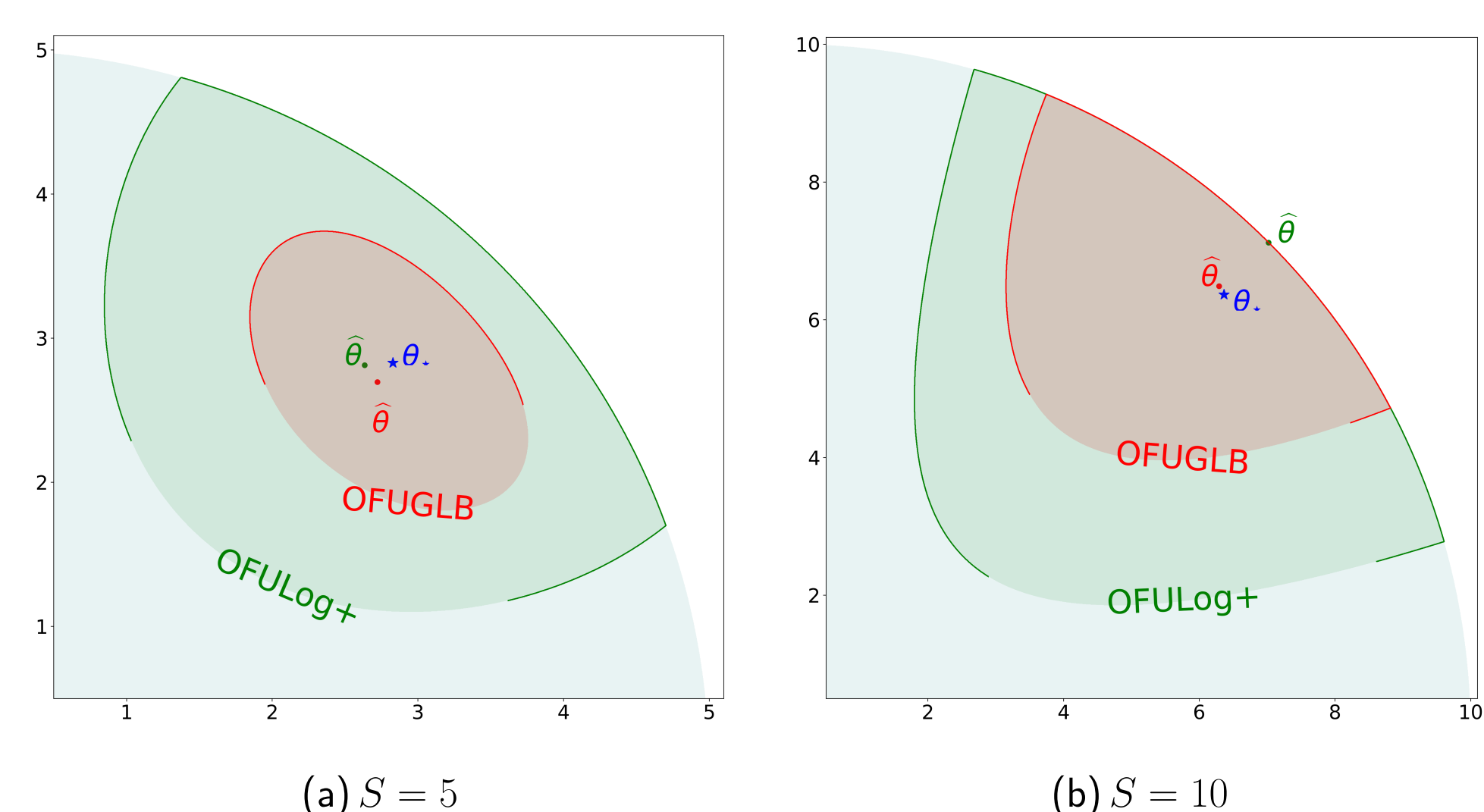


Figure 2: Confidence sets at $t = 4000$ from a single run.

Future Directions

- Extension to kernelized/functional GLMs?
- Implications to RLHF; see e.g., Das et al. [2024].
- Arm-set geometry-dependent transient term for GLBs
- Regret lower bound of general GLBs

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